

Principle of Minimum Renyi Correlation: from Marginals to Joint Distribution

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- Based on a joint work with



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U of Washington



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Stanford

Introduction

- **Example:** Genome-Wide Association Studies (GWAS)
 - ▶ SNPs $X_i \in \{0, 1, 2\}$

	X_1	X_2	$X_{3 \times 10^6}$
Indiv. 1	1	2		2
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- **Question:** How to model this data by a joint distribution?

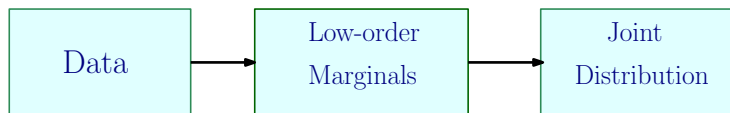
Introduction

- Not enough data to estimate ground-truth distribution

SNP sequences: $3^{3,000,000} \approx 10^{1,400,000}$
atoms in the universe $\approx 4 \times 10^{81}$

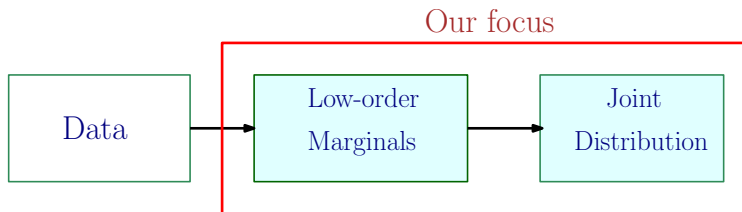
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- Low order marginals characterize a class of joint distributions \mathcal{C}



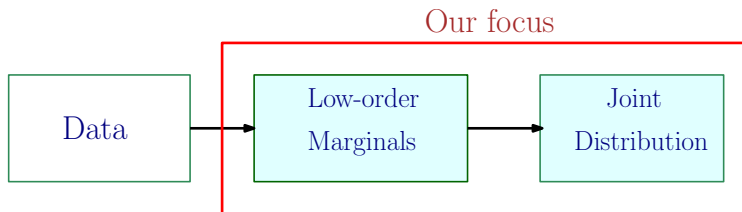
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- **Principle of maximum entropy:** pick the distribution maximizing *Shannon entropy* as a measure of *uncertainty*,

$$\operatorname{argmax}_{\mathbb{P} \in \mathcal{C}} H(\mathbb{P})$$

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 - ▶ SNPs X_i 's, trait Y

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- More sensible to minimize \mathbf{X} , Y dependence

$$\operatorname{argmin}_{\mathbb{P}_{\mathbf{X}, Y} \in \mathcal{C}} D_{\mathbb{P}}(\mathbf{X}; Y)$$

Principle of minimum Renyi correlation

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- Analytic structure of the minimizer
- Computation of the minimizing distribution

Minimum Renyi correlation distribution: continuous setting

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 - ▶ Second order moment, $\mathbb{E}[(\mathbf{X} \ Y)^T (\mathbf{X} \ Y)] = \boldsymbol{\Lambda} \in \mathbb{R}^{(p+1) \times (p+1)}$

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Theorem 1

Jointly Gaussian minimizes Renyi correlation.

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Theorem 1

Jointly Gaussian minimizes Renyi correlation.

Key reason: linearity of conditional expectation

$$\mathbb{E}[Y|X_1, \dots, X_p] = \sum_{i=0}^p c_i X_i$$

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Theorem 2

If there exists $\mathbb{P} \in \mathcal{C}$ with a separable conditional expectation,

$$\mathbb{E}_{\mathbb{P}} [Y \mid X_1, \dots, X_p] = \sum_i \gamma_i(X_i)$$

\mathbb{P} will minimize Renyi correlation.

Renyi minimizer distribution computation: Recipe

- Define \mathbf{X}_I as vector of indicator variables w.r.t. \mathbf{X} :

$$\mathbf{X}_{Imi+j} = \begin{cases} 1 & \text{if } \mathbf{X}_i = j \\ 0 & \text{otherwise} \end{cases}$$

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	X_1	Y
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Indiv. n	0		1

	X_{I_1}	X_{I_2}	X_{I_3}	Y
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- Find minimizer \mathbf{z}^* (*linear regression* on indicator variables):

$$\mathbf{z}^* \in \operatorname{argmin}_{\mathbf{z}} \mathbb{E} \left[(\mathbf{X}_I^T \mathbf{z} - Y)^2 \right] = \operatorname{argmin}_{\mathbf{z}} \mathbf{z}^T \mathbf{Q} \mathbf{z} + \mathbf{f}^T \mathbf{z} + 1$$

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- Define h as

$$h(X_1, \dots, X_p) = \frac{1}{2} (1 + \mathbf{z}^{*T} \mathbf{X}_I)$$

- If \mathbb{P} with separable conditional expectation exists

$$\mathbb{P}(Y = 1 | X_1, \dots, X_p) = h(X_1, \dots, X_p)$$

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Observation: necessary condition for existence of separable \mathbb{P} :

$$\forall x_1, \dots, x_p : 0 \leq h(x_1, \dots, x_p) \leq 1$$

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Question: How to check whether this condition holds?

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Theorem 3

Under the marginals consistency assumption, separable \mathbb{P} exists if and only if for a minimizer \mathbf{z}^* of

$$\min_{\mathbf{z}} \mathbb{E} \left[(\mathbf{X}_I^T \mathbf{z} - Y)^2 \right]$$

the separable function

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satisfies

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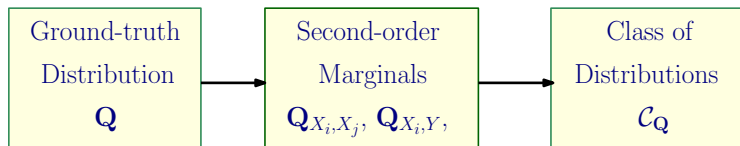
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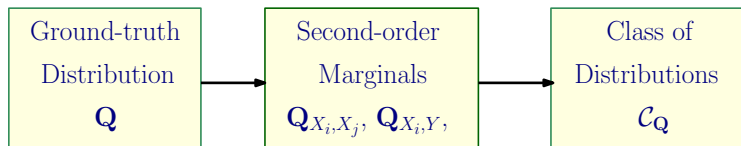
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Since h is separable, this condition can be checked in $O(mp)$.

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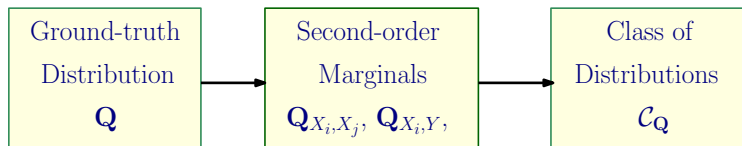


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Theorem 4

For \mathbb{P} uniform, there is an $\epsilon > 0$ such that for any Q in the ϵ -distance from \mathbb{P} , C_Q contains a distribution with separable conditional expectation.

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$$\max_{|S| \leq k} \mathcal{R}(\mathbf{X}_S, Y)$$

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$$\mathcal{R}(\mathbf{X}_S, Y) \geq \sqrt{1 - \min_{\mathbf{z}} \mathbb{E} [(\mathbf{X}_{I,S}^T \mathbf{z} - Y)^2]}$$

Tight under the additive structure assumption.

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- Using empirical average, equivalent to

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- Justification of *group lasso* for feature selection

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- There exists a certain separable structure in discrete minimizing distributions for given first and second order marginals
- We can compute conditional minimizing distribution by solving a linear regression problem.
- Principle of minimum Renyi correlation provides an interpretation for group LASSO as a variable selection method

Any Questions?