Exact Minimax Estimation for Phase Synchronization

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Anderson Zhang

Model $Y_{jk} = z_j \bar{z}_k + \sigma W_{jk} \in \mathbb{C}$ $1 \le j < k \le n$

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Parameter $z_j \in \mathbb{C}_1 = \{z \in \mathbb{C} : |z| = 1\}$ space

Model
$$Y_{jk} = z_j \overline{z}_k + \sigma W_{jk} \in \mathbb{C}$$
 $1 \le j < k \le n$
Im (W_{jk}) , Re $(W_{jk}) \sim N(0, 1/2)$

Parameter space

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Loss function

$$\ell(\widehat{z}, z) = \min_{a \in \mathbb{C}_1} \frac{1}{n} \sum_{j=1}^n |\widehat{z}_j a - z_j|^2$$



 $\max_{z \in \mathbb{C}_1^n} z^{\mathrm{H}} Y z$



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GPM

$$z_{j}^{(t)} = \frac{\sum_{k \in [n] \setminus \{j\}} Y_{jk} z_{k}^{(t-1)}}{\left| \sum_{k \in [n] \setminus \{j\}} Y_{jk} z_{k}^{(t-1)} \right|}$$



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Algorithms

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GPM

SDP

 $\max_{Z=Z^{\mathrm{H}}\in\mathbb{R}^{n\times n}}\operatorname{Tr}(YZ) \quad \text{subject to } \operatorname{diag}(Z)=I_{n} \text{ and } Z\succeq 0$



GPM



[Bandeira, Boumal & Singer 17] $\ell(\widehat{z}_{\text{MLE}}, z^*) \le C \frac{\sigma^2}{n}$ MLE



GPM



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GPM

[Bandeira, Boumal & Singer 17] $\hat{Z}_{SDP} = \hat{z}_{MLE} \hat{z}_{MLE}^{H}$ when $\sigma^2 = O(n^{1/2})$ **SDP**

MLE [Bandeira, Boumal & Singer 17] $\ell(\hat{z}_{\text{MLE}}, z^*) \le C \frac{\sigma^2}{n}$

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$$\ell(\widehat{z}, z) = \min_{a \in \mathbb{C}_1} \frac{1}{n} \sum_{j=1}^n |\widehat{z}_j a - z_j|^2$$

Theorem [G & Zhang]. Assume $\sigma^2 = o(np)$ and $\frac{np}{\log n} \to \infty$. Then,

 $\inf_{\widehat{z}\in\mathbb{C}_1^n}\sup_{z\in\mathbb{C}_1^n}\mathbb{E}_z\ell(\widehat{z},z)\geq (1-o(1))\frac{\sigma^2}{2n\nu}$

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Moreover, the MLE, GPM initialized by the leading eigenvector of $A \circ Y$, and the leading eigenvector of SDP all achieve $\ell(\hat{z}, z) \leq (1 + o(1)) \frac{\sigma^2}{2nn}$

with high probability.

Theorem [G & Zhang]. Assume $\sigma^2 = o(np)$ and $\frac{np}{\log n} \to \infty$. Then, $\inf_{\widehat{Z}} \sup_{z \in \mathbb{C}_1^n} \mathbb{E}_z \frac{1}{n^2} \|\widehat{Z} - zz^{H}\|_{F}^2 \ge (1 - o(1)) \frac{\sigma^2}{np}$.

Theorem [G & Zhang]. Assume $\sigma^2 = o(np)$ and $\frac{np}{\log n} \to \infty$. Then, $\inf_{\widehat{Z}} \sup_{z \in \mathbb{C}_1^n} \mathbb{E}_z \frac{1}{n^2} \|\widehat{Z} - zz^{\mathsf{H}}\|_{\mathsf{F}}^2 \ge (1 - o(1)) \frac{\sigma^2}{nn}$ Moreover, SDP achieves $\frac{1}{n^2} \|\widehat{Z} - zz^{\mathrm{H}}\|_{\mathrm{F}}^2 \le (1 + o(1)) \frac{\sigma^2}{nn}$ with high probability.

How to prove the results?

Statistical and Computational Guarantees of Lloyd's Algorithm and Its Variants

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December 8, 2016

Abstract

Clustering is a fundamental problem in statistics and machine learning. Lloyd's algorithm, proposed in 1957, is still possibly the most widely used clustering algorithm in practice due to its simplicity and empirical performance. However, there has been little theoretical investigation on the statistical and computational guarantees of Lloyd's algorithm. This paper is an attempt to bridge this gap between practice and theory. We investigate the performance of Lloyd's algorithm on clustering sub-Gaussian mixtures. Under an appropriate initialization for labels or centers, we show that Lloyd's algorithm converges to an exponentially small clustering error after an order of $\log n$ iterations, where n is the sample size. The error rate is shown to be minimax optimal. For the two-mixture case, we only require the initializer to be slightly better than random guess.

In addition, we extend the Lloyd's algorithm and its analysis to community detection and crowdsourcing, two problems that have received a lot of attention recently in statistics and machine learning. Two variants of Lloyd's algorithm are proposed respectively for community detection and crowdsourcing. On the theoretical side, we provide statistical and computational guarantees of the two algorithms, and the results improve upon some previous signal-to-noise ratio conditions in literature for both problems. Experimental results on simulated and real data sets demonstrate competitive performance of our algorithms to the state-of-the-art methods.

1 Introduction

Lloyd's algorithm, proposed in 1957 by Stuart Lloyd at Bell Labs [40], is still one of the most popular clustering algorithms used by practitioners, with a wide range of applications from computer vision [3], to astronomy [45] and to biology [26]. Although considerable innovations have been made on developing new provable and efficient clustering algorithms in the past six decades, Lloyd's algorithm has been consistently listed as one of the top ten data mining algorithms in several recent surveys [55].

$Y = \mathscr{X}_z(B) + w$

structure

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 $\begin{array}{l} \text{structure} \\ Y = \mathscr{X}_z(B) + w \end{array} \end{array}$ linear operator





Clustering

 $Y_i \sim N(\theta_{z_i}, I_d)$

Ranking $Y_{ij} \sim N(\beta(z_i - z_j), 1)$

Regression $Y_i \sim N(X_i^T \beta, 1)$

Clustering

$$Y_i \sim N(\theta_{z_i}, I_d)$$
$$\mathscr{X}_z : [\theta_1, ..., \theta_k] \mapsto [\theta_{z_1}, ..., \theta_{z_n}]$$

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$$Y_{ij} \sim N(\beta(z_i - z_j), 1)$$

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$$Y_{ij} \sim N(\beta(z_i - z_j), 1)$$

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Regression

 $Y_i \sim N(X_i^T \beta, 1)$ $\mathscr{X}_z : \beta \mapsto X\beta$
some local statistic T_i

 T_j

 $\mathbb{E}T_j = \mu_j(B, z_j)$

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 $\widehat{z}_j = \underset{a}{\operatorname{argmin}} \|T_j - \mu_j(B, a)\|^2$

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 $\widehat{z}_j = \underset{a}{\operatorname{argmin}} \|T_j - \mu_j(B, a)\|^2$ unknown, but $B \in \mathcal{B}_z$

 $\mathbb{E}T_j = \mu_j(B, z_j)$

 $z_{i}^{(t)} = \operatorname{argmin} ||T_{j} - \mu_{j}(\widehat{B}(z^{(t-1)}), a)||^{2}$ $\mathbf{\Omega}$

$$\mathbb{E}T_j = \mu_j(B, z_j)$$

$$z_j^{(t)} = \underset{a}{\operatorname{argmin}} \|T_j - \mu_j(\widehat{B}(z^{(t-1)}), a)\|^2$$

$$\widehat{B}(z) = \underset{B \in \mathcal{B}_z}{\operatorname{argmin}} \|Y - \mathscr{X}_z(B)\|^2$$

$$\mathbb{E}T_j = \mu_j(B, z_j)$$

$$\begin{cases} z_j^{(t)} = \underset{a}{\operatorname{argmin}} \|T_j - \mu_j(\widehat{B}(z^{(t-1)}), a)\|^2 \\ \widehat{B}(z) = \underset{B \in \mathcal{B}_z}{\operatorname{argmin}} \|Y - \mathscr{X}_z(B)\|^2 \end{cases}$$

loss function

$$\ell(z, z^*) = \sum_{j} \|\mu_j(B^*, z_j) - \mu_j(B^*, z_j^*)\|^2$$

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$$\geq \Delta_{\min}^2 \sum_{j} \mathbf{1}_{\{z_j \neq z_j^*\}}$$

loss function

$$\ell(z, z^*) = \sum_{j} \|\mu_j(B^*, z_j) - \mu_j(B^*, z_j^*)\|^2$$

error conditions

$$\frac{\operatorname{diff}\left(\mu_j(B(z), a), \mu_j(B(z^*), a)\right)}{\ell(z, z^*)} = o_{\mathbb{P}}(1)$$

diff
$$\left(\mu_j(\widehat{B}(z^*), a), \mu_j(B^*, a)\right) = o_{\mathbb{P}}(1)$$

loss function

$$\ell(z, z^*) = \sum_{j} \|\mu_j(B^*, z_j) - \mu_j(B^*, z_j^*)\|^2$$

error conditions

$$\max_{\{z:\ell(z,z^*)\leq\tau\}} \frac{\operatorname{diff}\left(\mu_j(B(z),a),\mu_j(B(z^*),a)\right)}{\ell(z,z^*)} = o_{\mathbb{P}}(1)$$

diff
$$\left(\mu_j(\widehat{B}(z^*), a), \mu_j(B^*, a)\right) = o_{\mathbb{P}}(1)$$

Convergence

Theorem [G & Zhang]. Assume

 $\ell(z^{(0)}, z^*) \le \tau$

and the conditions hold with the same τ .

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$$\ell(z^{(t)}, z^*) \le [\text{ideal error}] + \frac{1}{2}\ell(z^{(t-1)}, z^*)$$

or all $t \ge 1$.

$$\mathbb{E}T_j = \mu_j(B, z_j)$$

$$\begin{cases} z_j^{(t)} = \underset{a}{\operatorname{argmin}} \|T_j - \mu_j(\widehat{B}(z^{(t-1)}), a)\|^2 \\ \widehat{B}(z) = \underset{B \in \mathcal{B}_z}{\operatorname{argmin}} \|Y - \mathscr{X}_z(B)\|^2 \end{cases}$$

model

 $Y_i \sim N(\theta_{z_i}, I_d)$

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specialization

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algorithm

$$\begin{cases} z_i^{(t)} = \underset{a \in [k]}{\operatorname{argmin}} \|Y_i - \widehat{\theta}_a(z^{(t-1)})\|^2 \\\\ \widehat{\theta}_a(z) = \frac{\sum_{i=1}^n \mathbf{1}_{\{z_i = a\}} Y_i}{\sum_{i=1}^n \mathbf{1}_{\{z_i = a\}}} \end{cases}$$

model

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$$\begin{cases} z_i^{(t)} = \underset{a \in [k]}{\operatorname{argmin}} \|Y_i - \widehat{\theta}_a(z^{(t-1)})\|^2 \\\\ \widehat{\theta}_a(z) = \frac{\sum_{i=1}^n \mathbf{1}_{\{z_i = a\}} Y_i}{\sum_{i=1}^n \mathbf{1}_{\{z_i = a\}}} \\ \text{[Lloyd 57]} \end{cases}$$

Theorem and

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 $\frac{\min_{a \neq b} \|\theta_a^* - \theta_b^*\|^2}{k^2 (kd/n + 1)} \to \infty$

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Initialize by spectral clustering, and then

$$\frac{1}{n}\sum_{i=1}^{n} \mathbf{1}_{\left\{z_{i}^{(t)}\neq z_{i}^{*}\right\}} \leq \exp\left(-(1+o(1))\frac{\min_{a\neq b}\|\theta_{a}^{*}-\theta_{b}^{*}\|^{2}}{8}\right) + 2^{-t}$$

For all $t > 1$ with high probability.

TheoremAssume k = o(n),and $\frac{\min_{a \neq b} \|\theta_a^* - \theta_b^*\|^2}{k^2(kd/n+1)} \to \infty$ Initialize by spectral clustering, and then

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minimax rate

Theorem [Lu & Zhou 16]. Assume k = o(n), and $\frac{\min_{a \neq b} \|\theta_a^* - \theta_b^*\|^2}{k^2(kd/n + 1)} \to \infty$

Initialize by spectral clustering, and then

 $\frac{1}{n} \sum_{i=1}^{n} \mathbf{1}_{\left\{z_{i}^{(t)} \neq z_{i}^{*}\right\}} \leq \exp\left(-(1+o(1))\frac{\min_{a \neq b} \|\theta_{a}^{*} - \theta_{b}^{*}\|^{2}}{8}\right) + 2^{-t}$ for all $t \geq 1$ with high probability.

minimax rate

$$Y = zz^{\mathrm{H}} + \sigma W$$

 $Y = zz^{\rm H} + \sigma W$ quadratic





$$Y = z\beta^{\rm H} + \sigma W$$

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linear in z (instead of quadratic)
$$Y = z\beta^{\mathrm{H}} + \sigma W \qquad \mathscr{X}_{z} : \beta \mapsto z\beta^{\mathrm{H}}$$
$$\beta \in \mathcal{B}_{z} = \{z\}$$

linear in z (instead of quadratic)

specialization $T_i = Y_i$

 $Y = z\beta^{H} + \sigma W$ $\mathscr{X}_{z}: \beta \mapsto z\beta^{H}$ $\beta \in \mathcal{B}_{z} = \{z\}$

linear in z (instead of quadratic)

specialization
$$T_i = Y_i$$
 $\mu(\beta, a) = a\beta$

 $Y = z\beta^{H} + \sigma W$ $\mathscr{X}_{z}: \beta \mapsto z\beta^{H}$ $\beta \in \mathcal{B}_{z} = \{z\}$

linear in z (instead of quadratic)

$$\begin{cases} \beta^{(t)} = \underset{\beta \in \{z^{(t-1)}\}}{\operatorname{argmin}} \|Y - z^{(t-1)}\beta^{\mathsf{H}}\|_{\mathsf{F}}^{2} \\ z_{j}^{(t)} = \underset{a \in \mathbb{C}_{1}}{\operatorname{argmin}} \|Y_{j} - a\beta^{(t)}\|^{2} \end{cases}$$

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linear in z (instead of quadratic)

$$z_{j}^{(t)} = \frac{\sum_{k \in [n] \setminus \{j\}} Y_{jk} z_{k}^{(t-1)}}{\left| \sum_{k \in [n] \setminus \{j\}} Y_{jk} z_{k}^{(t-1)} \right|}$$

 $Y = z\beta^{H} + \sigma W$ $\mathscr{X}_{z}: \beta \mapsto z\beta^{H}$ $\beta \in \mathcal{B}_{z} = \{z\}$

linear in z (instead of quadratic)

$$z_i^{(t)} = \frac{[(A \circ Y)z^{(t-1)}]_i}{|[(A \circ Y)z^{(t-1)}]_i|}$$



$$z^{(t)} = f(z^{(t-1)}) \quad \checkmark \quad z_i^{(t)} = \frac{[(A \circ Y)z^{(t-1)}]_i}{|[(A \circ Y)z^{(t-1)}]_i|}$$

Lemma. Assume $\sigma^2 = o(np)$ and $p \gg \frac{\log n}{n}$. For any $\gamma = o(1)$, we have $\mathbb{P}\left(\ell(f(z), z^*) \le \delta\ell(z, z^*) + (1 + \delta)\frac{\sigma^2}{2np} \text{ for any } z \in \mathbb{C}_1^n \text{ s.t. } \ell(z, z^*) \le \gamma\right) \ge 1 - \delta$ for some $\delta = o(1)$.

Corollary. Assume $\sigma^2 = o(np)$ and $p \gg \frac{\log n}{n}$. Initialized by PCA, the power method satisfies

with high probability after $\log\left(\frac{np}{\sigma^2}\right)$ iterations.

 $\ell(\widehat{z}, z) \le (1 + o(1)) \frac{\sigma^2}{2nn}$

$\mathsf{MLE} \quad \max_{z \in \mathbb{C}_1^n} z^{\mathrm{H}} (A \circ Y) z$





Corollary. Assume $\sigma^2 = o(np)$ and $p \gg \frac{\log n}{n}$. The MLE satisfies $\ell(\hat{z}, z) \leq (1 + o(1)) \frac{\sigma^2}{2np}$ with high probability.

MLE $\max \operatorname{Tr}((A \circ Y)zz^{\mathrm{H}})$ s.t. $|z_j| = 1$ for all $j \in [n]$

MLE
$$\max \operatorname{Tr}((A \circ Y)zz^{H})$$

s.t. $|z_j| = 1$ for all $j \in [n]$

GPM
$$z_{j}^{(t)} = \frac{\sum_{k \in [n] \setminus \{j\}} A_{jk} Y_{jk} z_{k}^{(t-1)}}{\left| \sum_{k \in [n] \setminus \{j\}} A_{jk} Y_{jk} z_{k}^{(t-1)} \right|}$$





SDP
$$\max \operatorname{Tr}((A \circ Y)Z)$$

s.t. diag(Z) = I_n and $Z \succeq 0$



SDP
s.t. diag
$$(Z) = I_n$$
 and $Z \succeq 0$
 $Z = V^H V$



SDP
s.t. diag(Z) =
$$I_n$$
 and $Z \succeq 0$

$$\sum_{V \in \mathbb{C}^{n \times n}} \operatorname{Tr}((A \circ Y)V^{\mathrm{H}}V)$$
s.t. $\|V_j\|^2 = 1$ for all $j \in [n]$



SDP
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$$\operatorname{diag}(Z) = I_n \text{ and } Z \succeq 0$$

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s.t. $\|V_j\|^2 = 1$ for all $j \in [n]$

[Burer & Monteiro 03]



SDP
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$$\sum_{V \in \mathbb{C}^{n \times n}} \operatorname{Tr}((A \circ Y)V^{\mathrm{H}}V)$$
s.t. $\|V_j\|^2 = 1$ for all $j \in [n]$

$$V_{j}^{(t)} = \frac{\sum_{k \in [n] \setminus \{j\}} A_{jk} \bar{Y}_{jk} V_{k}^{(t-1)}}{\left\| \sum_{k \in [n] \setminus \{j\}} A_{jk} \bar{Y}_{jk} V_{k}^{(t-1)} \right\|}$$

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$$V^{(t)} = f(V^{(t-1)}) \quad \checkmark \quad V_{j}^{(t)} = \frac{\sum_{k \in [n] \setminus \{j\}} A_{jk} \bar{Y}_{jk} V_{k}^{(t-1)}}{\left\| \sum_{k \in [n] \setminus \{j\}} A_{jk} \bar{Y}_{jk} V_{k}^{(t-1)} \right\|}$$

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Lemma. Assume $\sigma^2 = o(np)$ and $p \gg \frac{\log n}{n}$. For any $\gamma = o(1)$, we have $\mathbb{P}\left(\ell(f(V), z^*) \le \delta\ell(V, z^*) + (1 + \delta)\frac{\sigma^2}{2np} \text{ for any } z \in \mathbb{C}_1^n \text{ s.t. } \ell(V, z^*) \le \gamma\right) \ge 1 - \delta$ for some $\delta = o(1)$.

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for some $\delta = o(1)$.

$$\ell(\widehat{V}, z^*) = \min_{a \in \mathbb{C}^n : ||a||^2 = 1} \frac{1}{n} \sum_{j=1}^n ||\widehat{V}_j - \overline{z}_j^* a||^2$$

 $\max \mathsf{Tr}((A \circ Y)Z)$ SDP s.t. diag(Z) = I_n and $Z \succeq 0$



SDP

$$\begin{array}{l} \max \operatorname{Tr}((A \circ Y)Z) \\ \text{s.t. } \operatorname{diag}(Z) = I_n \text{ and } Z \succeq 0 \end{array} \qquad \widehat{V} = \widehat{V}^{\widehat{V}} \widehat{V} = \widehat{f}(\widehat{V}) \\
\end{array}$$
Corollary. Assume $\sigma^2 = o(np)$ and $p \gg \frac{\log n}{n}$.
The SDP satisfies
 $\ell(\widehat{V}, z) \leq (1 + o(1)) \frac{\sigma^2}{2np}, \\ \frac{1}{n^2} \|\widehat{Z} - zz^{\mathsf{H}}\|_{\mathrm{F}}^2 \leq (1 + o(1)) \frac{\sigma^2}{np}, \\ \ell(\widehat{z}, z) \leq (1 + o(1)) \frac{\sigma^2}{2np} \\
\end{array}$
with high probability.



proof of lower bound

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 $\inf_{\widehat{z}\in\mathbb{C}_1^n}\sup_{z\in\mathbb{C}_1^n}\mathbb{E}_z\min_{\theta\in\mathbb{R}}\sum_{j=1}^n|\widehat{z}_je^{i\theta}-z_j|^2$



 $\inf_{\widehat{z} \in \mathbb{C}_{1}^{n}} \sup_{z \in \mathbb{C}_{1}^{n}} \mathbb{E}_{z} \min_{\theta \in \mathbb{R}} \sum_{j=1}^{n} |\widehat{z}_{j}e^{i\theta} - z_{j}|^{2}$ not separable


$$\begin{aligned} \inf_{\widehat{z} \in \mathbb{C}_{1}^{n}} \sup_{z \in \mathbb{C}_{1}^{n}} \mathbb{E}_{z} \min_{\theta \in \mathbb{R}} \sum_{j=1}^{n} |\widehat{z}_{j}e^{i\theta} - \underline{z}_{j}|^{2} & \text{not separable} \\ \geq & \frac{1}{2n} \inf_{\widehat{z} \in \mathbb{C}_{1}^{n}} \sup_{z \in \mathbb{C}_{1}^{n}} \mathbb{E}_{z} \|\widehat{z}\widehat{z}^{\mathrm{H}} - zz^{\mathrm{H}}\|_{\mathrm{F}}^{2} \end{aligned}$$







$$\geq \frac{1}{2n} \inf_{\widehat{z} \in \mathbb{C}_{1}^{n}} \sup_{z \in \mathbb{C}_{1}^{n}} \mathbb{E}_{z} \|\widehat{z}\widehat{z}^{\mathsf{H}} - zz^{\mathsf{H}}\|_{\mathrm{F}}^{2} \qquad \text{separable}$$

$$\geq \frac{1}{2n} \inf_{\widehat{z}} \sum_{1 \leq j \neq k \leq n} \left(\int \prod_{l=1}^{n} \pi(z_{l}) \right) \mathbb{E}_{z} |\widehat{z}_{j}\overline{\widehat{z}}_{k} - z_{j}\overline{z}_{k}|^{2} dz$$

$$\geq \frac{1}{2n} \sum_{1 \leq j \neq k \leq n} \mathbb{E}_{z_{-(j,k)} \sim \pi} \inf_{\widehat{T}} \int \pi(z_{j}) \pi(z_{k}) \mathbb{E}_{z} |\widehat{T} - z_{j}\overline{z}_{k}|^{2} dz_{j} dz_{k}$$





Model

 $Y_{jk} = z_j z_k + \sigma W_{jk}$

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Parameter space

 $z_j \in \{-1, 1\}$

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 $z_i \in \{-1, 1\}$

Missing data

 $A_{jk} \sim \text{Bernoulli}(p)$

MLE

 $\max_{z \in \{-1,1\}^n} z^{\mathrm{T}} (A \circ Y) z$

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GPM

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SDP $\max_{Z=Z^{\mathrm{T}}\in\mathbb{R}^{n\times n}} \operatorname{Tr}((A\circ Y)Z)$ subject to $\operatorname{diag}(Z) = I_n$ and $Z \succeq 0$

Theorem [G & Zhang]. Assume $\sigma^2 = o(np)$ and $\frac{np}{\log n} \to \infty$. Then,

$$\inf_{\widehat{z} \in \{-1,1\}^n} \sup_{z \in \{-1,1\}^n} \mathbb{E}_z \ell(\widehat{z}, z) \ge \exp\left(-(1+o(1))\frac{np}{2\sigma^2}\right)$$

Moreover, the MLE, GPM initialized by the leading eigenvector of $A \circ Y$, and the leading eigenvector of SDP all achieve

$$\ell(\widehat{z}, z) \le \exp\left(-(1 - o(1))\frac{np}{2\sigma^2}\right)$$

with high probability.

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"p=1" by [Fei & Chen 20]

phase sync

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s.t. $\operatorname{diag}(Z) = I_n$ and $Z \succeq 0$

phase sync

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Comparison

phase sync

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$$\max_{Z=Z^{\mathrm{H}}\in\mathbb{C}^{n\times n}} \operatorname{Tr}((A\circ Y)Z)$$

s.t. diag(Z) = I_n and $Z \succeq 0$

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s.t. diag(Z) = I_n and $Z \succ 0$

$$\exp\left(-(1-o(1))\frac{np}{2\sigma^2}\right)$$

For any $x, y \in \mathbb{C}^n$ such that ||y|| = 1 and $\operatorname{Re}(y^{\mathrm{H}}x) > 0$, we have $\left\|\frac{x}{\|x\|} - y\right\|^2 \le \frac{\|(I_n - yy^{\mathrm{H}})x\|^2 + |\operatorname{Im}(y^{\mathrm{H}}x)|^2}{|\operatorname{Re}(y^{\mathrm{H}}x)|^2}$

Comparison

phase sync

$$\max_{\substack{Z=Z^{\mathrm{H}}\in\mathbb{C}^{n\times n}}} \operatorname{Tr}((A\circ Y)Z)$$

s.t. diag(Z) = I_n and $Z \succeq 0$ $(1+o(1))\frac{\sigma^2}{2np}$

Z2 sync

$$\max_{\substack{Z=Z^{\mathrm{T}}\in\mathbb{R}^{n\times n}\\ \text{s.t. } \operatorname{diag}(Z)=I_{n} \text{ and } Z\succeq 0}} \operatorname{Tr}((A\circ Y)Z) \exp\left(-(1-o(1))\frac{np}{2\sigma^{2}}\right)$$

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$$x, y \in \mathbb{C}^n$$
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Model $Y_{ij} = Z_i Z_j^T + \sigma W_{ij} \in \mathbb{R}^{d \times d}$ $Z_i \in SO(d)$

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Model $Y_{ij} = Z_i Z_j^{\mathrm{T}} + \sigma W_{ij} \in \mathbb{R}^{d \times d}$ $Z_i \in \mathsf{SO}(d)$ $A_{jk} \sim \mathrm{Bernoulli}(p)$ **LOSS** $\ell(\widehat{Z}, Z) = \min_{U \in \mathsf{SO}(d)} \frac{1}{n} \sum_{j=1}^n \|\widehat{Z}_j U - Z_j\|_{\mathrm{F}}^2$

Theorem [G & Zhang]. Assume $\sigma^2 = o(np)$ and $\frac{np}{\log n} \to \infty$. Then, $\inf_{\widehat{Z} \in SO(d)^n | Z \in SO(d)^n} \mathbb{E}_n \ell(\widehat{Z}, Z) \ge (1 - o(1)) \frac{(d - 1)d\sigma^2}{2np}$. Moreover, both MLE and GPM achieve $\ell(\widehat{Z}, Z) \le (1 + o(1)) \frac{(d - 1)d\sigma^2}{2np}$ with high probability. Thank You