# Quantitative trait evolution with mutations of large effect

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Joshua G. Schraiber and Michael J. Landis Mutations of large effect

# Quantitative traits

- Traits that vary continuously in populations
  - Mass
  - Height
  - Bristle number (approx)
- $\cdot$  Adaption
  - Low oxygen tolerance
- Disease
  - Obesity
- · Agriculture
  - Fruit yield



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## Fisher's quantitative trait model

- n biallelic loci impacting trait
- · Allele *j* at locus *i* has effect *a<sub>ij</sub>*
- Phenotype is  $X_k = \sum_{i,j} a_{ij} z_{ijk}$ 
  - z<sub>ijk</sub> is number of copies of allele j at locus i in individual k
- $\cdot \mathbb{E}(\bar{X}) = 2n\mathbb{E}(p)\mathbb{E}(a)$
- $\cdot \mathbb{E}(\operatorname{Var}(X)) = 2n\mathbb{E}(p(1-p))\mathbb{E}(a^2)$
- Environmental variation induces extra variability



# Neutral genetic variance

- · Equilibrium genetic variance
  - $\mathbb{E}(V_A(\infty)) = 2NV_m$
- ·  $V_m$  is the additive variance of new mutations
  - $V_m pprox 2\mu\sigma^2$
  - $\mu$ : is trait-wide mutation rate
  - $\sigma^2$ : variance of mutant effect size distribution
- · Var( $V_A(\infty)$ )  $\approx 4N\mu\sigma^4$  (Lynch and Hill 1986)



## Lande's Brownian motion model

- · Approximation to full model
  - Don't keep track of individual loci
  - Assume genetic variance constant in time
  - Mutations have small effect sizes
- · Can incorporate selection via Ornstein-Uhlenbeck model
- Extremely influential in comparative biology
  - c.f. Felsenstein: independent contrasts



## Evidence for non-Brownian evolution



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- · Haploid population of size N
- · Trait governed by *n* loci
- $\cdot\,$  Each locus has an independent coalescent tree
- · Each locus has mutation rate  $\frac{\theta}{2}$  (coalescent time units)
- When a mutation happens, effect Y is drawn from distribution with density p(y).

## Example



## Mutational effects

 $\cdot$  Common assumption: many loci of very small effect





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## Large effect mutations

- · In some cases, large effect mutations may occur
  - Transcription factor binding sites
  - Null mutants upstream in pathways



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#### Characteristic function

For a random variable X drawn from the probability measure  $\mu(\cdot)$ , the function

$$egin{array}{rcl} \phi_X(k) &=& \mathbb{E}(e^{ikX}) \ &=& \int e^{ikx} d\mu(x) \end{array}$$

is called the characteristic function

#### Sums of random variables

If  $X_1, \ldots, X_n$  are independent random variables, and  $X = \sum_i X_i$  then

$$\phi_X(k) = \prod_i \phi_{X_i}(k)$$

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## Compound Poisson process

#### Definition

Let N(t) be a rate  $\lambda$  Poisson process, and  $(Y_i, i \ge 1)$  a sequence of independent and identically distributed random variables. Then

$$X(t) = \sum_{i=1}^{N(t)} Y_i$$

is called a compound Poisson process.



#### CF of a CP process

If X(t) is a rate  $\lambda$  Compound Poisson process, and  $\psi(k)$  is the characteristic function of the jump distribution, then the characteristic function of X is

$$\phi_t(k) = e^{\lambda t(\psi(k)-1)}$$



- $X_1$  and  $X_2$  are independent  $CP(\theta/2)$ processes run for time  $T_2$
- Joint distribution depends on root value
- Consider  $Z = X_2 X_1$ 
  - Doesn't depend on root



## One locus, sample of size 2 (CF)



- · The root of the tree must be specified
- A new individual could coalesce *more anciently* than the current root
  - For each *n*, there is a fixed root
  - Ensuring consistency could be hard

## The root problem



#### Normalization

If  $X = (X_0, \ldots, X_{n-1})$  are the trait values in the sample, then the random vector

$$Z = (Z_1, \dots, Z_{n-1}) = (X_1 - X_0, \dots, X_{n-1} - X_0)$$

does not depend on the root value

- $\cdot$  "Normalizing" by an arbitrary individual makes *n* unimportant
  - Intuition: now every sample has to follow the path through the tree to individual 0, so the root doesn't matter

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## Characteristic function for n = 3, and symmetric mutational effects

$$\phi_{Z_1,Z_2}(k_1,k_2) = \frac{\frac{1}{1+\theta(1-\psi(k_1))} + \frac{1}{1+\theta(1-\psi(k_2))} + \frac{1}{1+\theta(1-\psi(k_1+k_2))}}{3 - \frac{\theta}{2}(\psi(k_1) + \psi(k_2) + \psi(k_1+k_2) - 3)}$$

- · Sketch of derivation:
  - Condition on topology (each of 3 topologies equally likely)
  - Compute characteristic function for each topology while conditioning on coalescence times
  - Integrate over coalescence times
  - Take weighted sum of characteristic functions for each topology

# Sending the number of loci off to infinity

- $\cdot\,$  We would like some nice limit as the number of loci increases
  - Trivial result: "uniform on  $\mathbb{R}$ " or  $\delta(x)$
- $\cdot$  Need to decrease the effect size or mutation rate of each locus
- · Three nontrivial limits
  - Mutation rate per locus decreases but effect sizes remain constaint
  - Mutation rate per locus remains constant but effect sizes decrease and effect size distribution does not have fat tails
  - Mutation rate per locus remains constant but effect size decreases and effect size distribution has fat tail

#### Correlated CPP limit

As  $n \uparrow \infty$  and  $\theta \downarrow 0$  such that  $n\theta \rightarrow \Theta$ ,

$$\phi_{Z_1,Z_2}(k_1,k_2) \to e^{\frac{\Theta}{2}(\psi(k_1)+\psi(k_2)+\psi(k_1+k_2)-3)}$$

which is the characteristic function of two correlated compound Poisson processes.

- · Perhaps not very biologically relevant
- · Will ignore for rest of talk

#### Bivariate Gaussian limit

As  $n \uparrow \infty$  and the second moment of the effect distrbution,  $\tau^2$ ,  $\downarrow 0$  such that  $n\tau^2 \rightarrow \sigma^2$ ,

$$\phi_{Z_1,Z_2}(k_1,k_2) \to e^{-\frac{\theta}{2}\sigma^2(k_1^2+k_1k_2+k_2^2)}$$

which is the characteristic function of a Bivariate Gaussian distribution with mean vector (0,0) and variance-covariance matrix

$$\Sigma = \theta \sigma^2 \begin{bmatrix} 1 & 1/2 \\ 1/2 & 1 \end{bmatrix}$$

#### Bivariate stable distribution

Assume that  $p(y) \sim \kappa |y|^{-(\alpha+1)}$  and set  $t = \kappa \pi (\sin(\alpha \pi/2)\Gamma(\alpha)\alpha)^{-1}$ . As  $n \uparrow \infty$  and  $t \downarrow 0$  such that  $nt \rightarrow c$ ,

$$\phi_{Z_1,Z_2}(k_1,k_2) \to e^{-\frac{1}{2}\theta c(|k_1|^{\alpha}+|k_2|^{\alpha}+|k_1+k_2|^{\alpha})}$$

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where  $c = \tilde{c} \frac{\pi}{\sin(\frac{\alpha \pi}{2})}$ , which is the characteristic function of a bivariate  $\alpha$ -stable distribution

#### Theorem

If  $(X_1, X_2, ...)$  is an infinitely exchangeable random vector, then the probability density of  $(X_1 = x_1, ..., X_n = x_n)$  is a mixture over i.i.d. probability densities. That is,

$$p(x_1,\ldots,x_n) = \int \prod_i p_{\theta}(x_i) \nu(d\theta)$$

• Suggests that we can find a de Finetti measure such that all our normalized samples are i.i.d.

## A guess in the normal case

- · Given the mean, the trait is distributed  $\mathcal{N}(0, \frac{\theta}{2}\sigma^2)$
- · The mean is itself distributed  $\mathcal{N}(0, \frac{\theta}{2}\sigma^2)$
- · Law of total variance gives  $Var(X) = 2N\mu\sigma^2$ 
  - Same as classical derivation

$$\begin{split} \phi(k_1, k_2) &= \int \left( \int \prod_{l=1}^2 e^{ik_l x_l} \frac{1}{\sqrt{\pi \theta \sigma^2}} e^{-\frac{(m-x_l)^2}{\theta \sigma^2}} dx_l \right) \frac{1}{\sqrt{\pi \theta \sigma^2}} e^{-\frac{m^2}{\theta \sigma^2}} dm \\ &= e^{-\frac{\theta}{4} \sigma^2 \left(k_1^2 + k_2^2\right)} \int e^{im(k_1 + k_2)} \frac{1}{\sqrt{\pi \theta \sigma^2}} e^{-\frac{m^2}{\theta \sigma^2}} dm \\ &= e^{-\frac{\theta}{4} \sigma^2 \left(k_1^2 + k_2^2\right)} e^{-\frac{\theta}{4} \sigma^2 (k_1 + k_2)^2} \\ &= e^{-\frac{\theta}{2} \sigma^2 \left(k_1^2 + k_1 k_2 + k_2^2\right)} \end{split}$$

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## Tempting interpretation and conjecture

- $\cdot$  Conditioning on the amount of evolution "exclusive to sample 0"
  - Integrate over whether sample 1 or sample 2 coalesces with sample 0 first
- $\cdot$  Suggests that this can be extended to arbitrary sample sizes

Conjecture about larger samples (Gaussian limit)

When the mutation kernel has only small effect mutations, the limit distribution is normal with variance  $\frac{\theta}{2}\sigma^2$  and random mean.

# Testing the Gaussian conjecture

- $\cdot\,$  Simulate quantitative traits according to the coalescent model
- $\cdot$  Use KS test to assess convergence in the limit



#### · Bivariate case

- Independently lpha-stable with random median
- Analogous to Gaussian limit

## Conjecture about larger samples (Stable limit)

When the mutation kernel has large mutational effects, the limit distribution is  $\alpha$ -stable with scale  $\left(\frac{\theta}{2}c\right)^{1/\alpha}$  and random median.

## Testing the stable conjecture

- $\cdot\,$  Simulate quantitative traits according to the coalescent model
- $\cdot$  Use KS test to assess convergence in the limit





## Stable limit for sample of size 4

- · Tedious computation
  - Need to average over 18 trees.
- · Conjecture would require

$$\phi_{Z_1,Z_2,Z_3}(k_1,k_2,k_3) \to e^{-\frac{\theta}{2}c(|k_1|^{\alpha}+|k_2|^{\alpha}+|k_3|^{\alpha}+|k_1+k_2+k_3|^{\alpha})}$$

#### Trivariate characteristic function

For a sample of size three, under the same conditions as the bivariate limit,

$$\begin{split} \phi_{Z_1,Z_2,Z_3}(k_1,k_2,k_3) \to \exp\{-\frac{\theta}{3}c(|k_1|^{\alpha}+|k_2|^{\alpha}+|k_3|^{\alpha}\\ &+\frac{1}{2}|k_1+k_2|^{\alpha}+\frac{1}{2}|k_1+k_3|^{\alpha}+\frac{1}{2}|k_2+k_3|^{\alpha}\\ &+|k_1+k_2+k_3|^{\alpha})\} \end{split}$$

## Conjecture (Stable limit)

When the mutation kernel has large mutational effects, the limit distribution is  $\alpha\text{-stable}$  with random parameters

- $\cdot$  Not even sure that this is true
  - A mixture of stable distributions?

- · Neurospora crassa mRNA-seq
  - Ellison et al (2011), PNAS
  - Restrict to genes with  $\mathsf{FPKM}>1$
- $\cdot\,$  For each gene fit normal distribution,  $\alpha\text{-stable}$  distribution
- · Bootstrap likelihood ratio test ( $H_0: \alpha = 2$  vs.  $H_1: \alpha < 2$ )

- · Preliminary!
  - Only analyzed 346 genes

## Distribution of p-values



• 54 genes significant at FDR of 32%

## Distribution of $\alpha$ in significant genes



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## Cherry-picked examples





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NCU02435

NCU07659





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## Conclusions

- $\cdot$  Introduced a coalescent model of quantitative trait evolution
  - Provided exact formulas for the characteristic functions for samples of size 2, 3, 4
  - Found limiting distributions as the number of loci became large in each case
  - Conjectured about extending these limits to larger samples
- Extensions
  - Samples not contemporaneous
  - Population structure
  - Diploids
- · Why a neutral model?
  - Analytically tractable calculations
  - Intuition for what to expect with weak selection
  - Null model